

### Corollary

Suppose  $N$  is a PRM on  $(S, \mathcal{F}, \eta)$ . Then for each  $A \in \mathcal{F}$ ,  $N(\cdot \cap A)$  is a PRM on  $(A, A \cap \mathcal{F}, \eta(\cdot; A))$

Further, if  $A, B \in \mathcal{F}$  and  $A \cap B = \emptyset$ , then  $N(\cdot \cap A) \perp\!\!\!\perp N(\cdot \cap B)$ .

Corollary. Suppose  $N$  is a PRM on  $(S, \mathcal{F}, \eta)$

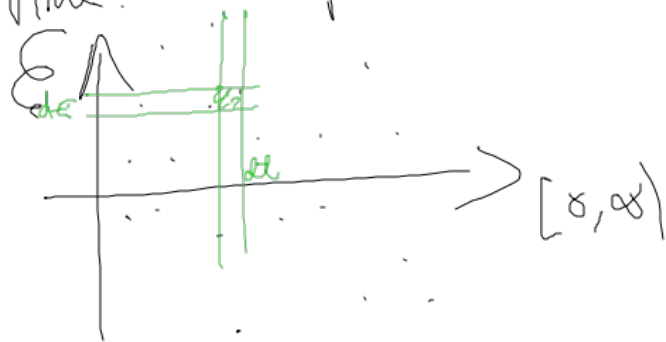
Then the support of  $N$  is P.a.s. countable.

If in addition  $\eta$  is a finite measure, then the support of  $N$  is a.s. finite (in fact it consists of a Poisson distributed # of pts.)

Remark: If  $\eta$  has atoms i.e.  $\exists s \in S$  s.t.  $\{s\} \in \mathcal{F}$  and  $\eta(\{s\}) > 0$ , then we can get a "pile-up" of points at  $\{s\}$ . If  $\eta$  is diffuse (no atoms) then the probability that  $N(\{s\}) > 1$  is zero. If  $s \in S$  such that  $\{s\} \in \mathcal{F}$ .

# Poisson Point Processes

In the case that  $S = [0, \infty) \times E$  for some space  $E$ , we can think of points in the support of an associated PRM as arriving in time.



The support of an associated PRM on  $S = [0, \infty) \times E$  is called a Poisson Point Process if the intensity measure  $\eta(dt \times dE) = \underline{dt \times n(dE)}$  for some  $\sigma$ -finite measure  $n$ .  
diffuse!

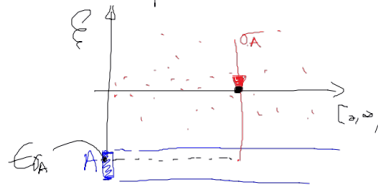
Intuitively: A "point" arrives in PPP at time  $t$  and "pos"  $\in E$  with prob  $dt n(dE) + r(dt)$



## Thinning Theorem for PPP

Notation Given measure (configuration)  $\mu$  on  $\mathcal{E}$   
 $(t, \epsilon_t)$  are the points supporting the PRM on  $[0, \infty) \times \mathcal{E}$  with intensity  $dt \times d\mu$   
 If we define  $\epsilon_t = \emptyset$  for  $t$  such that  $t$  is not a "time" in the support of our PRM  
 then we can think of  $\{\epsilon_t : t \geq 0\}$  as our PPP on  $\mathcal{E}$ .

Theorem, Suppose  $A \in \mathcal{E}$  is an  $n$ -measurable set such that  $n(A) < \infty$ . Let  
 $\sigma_A = \inf\{t > 0 : \epsilon_t \in A\}$ .



(i)  $\sigma_A$  is a stopping time w.r.t. the natural filtration of the underlying PPP  
 $\mathcal{F}_t = \sigma\left(N([0, s] \times B) = k : \begin{matrix} s \leq t \\ B \in \mathcal{E} / (k \in \mathbb{Z}, l \in \mathbb{Z}) \end{matrix}\right)$   
*sigma-algebra of E*

Moreover  $\sigma_A \sim \exp(n(A))$ .

(ii)  $\epsilon_{\sigma_A} \perp \{\epsilon_t : t < \sigma_A\}$  and has law given by  $\frac{n(\cdot; A)}{n(A)}$

(iii)  $\{\epsilon_t : t < \sigma_A\}$  has the law of a PPP with intensity  $dt \times n(d\epsilon; \mathcal{E} \setminus A)$  which is stopped at an independent and exponentially distributed random time with parameter  $n(A)$ .

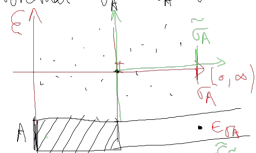
# (Sketch #)

Suppose  $f: [0, \infty) \times E \rightarrow [0, \infty)$  is bounded and measurable w.r.t.  $dt \times d\epsilon$ .

$\mathbb{E} \left[ e^{-\int_0^{\sigma_A} f(s, \epsilon) N(ds \times d\epsilon)} ; \sigma_A > t, \epsilon_{\sigma_A} \in B \right]$   
 $B \subseteq A$  and  $B$  is  $n$ -measurable.  $\Rightarrow \sigma_A$  is a stopping time.

where  $\{\tilde{\epsilon}_s : s > 0\} = \{\epsilon_{t+s} : s > 0\}$   
 $\tilde{\sigma}_A = \inf \{s > 0 : \tilde{\epsilon}_s \in A\}$ .

Note that  $\tilde{\sigma}_A = \sigma_A - t$  on  $\{\sigma_A > t\}$ .



$\mathbb{E} \left( e^{-\int_0^t \int_{\Sigma \setminus A} f(s, \epsilon) N(ds \times d\epsilon)} ; N([0, t] \times A) = 0, \tilde{\epsilon}_{\tilde{\sigma}_A} \in B \right)$

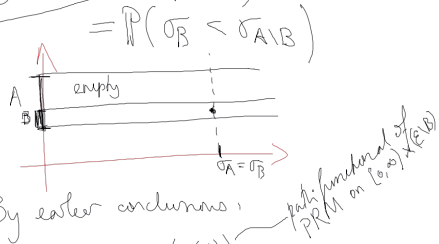
$= \mathbb{E} \left( e^{-\int_0^t \int_{\Sigma \setminus A} f(s, \epsilon) N(ds \times d\epsilon)} \right) \times \mathbb{P}(N([0, t] \times A) = 0)$   
 $\times \mathbb{P}(\tilde{\epsilon}_{\tilde{\sigma}_A} \in B)$

$= \mathbb{E} \left( e^{-\int_0^t \int_{\Sigma \setminus A} f(s, \epsilon) N(ds \times d\epsilon)} \right) \times e^{-t \eta(A)} \times \mathbb{P}(\epsilon_{\sigma_A} \in B)$

This proposition tells us:  
 (i)  $\sigma_A \sim \exp(\eta(A))$   
 (ii)  $\sigma_A \perp \{ \epsilon_t : t < \sigma_A \}$   
 (iii)  $\{ \epsilon_t : t < \sigma_A \} \sim \text{PPP}(dt \times \eta(d\epsilon; A))$   
 stopped at  $\exp(\eta(A))$ -time.

Finally note that  $\mathbb{P}(\epsilon_{\sigma_A} \in B)$

$= \mathbb{P}(\sigma_B < \sigma_{A \setminus B})$



By earlier conclusions,

$\sigma_{A \setminus B} \sim \exp(\eta(A \setminus B))$   
 $\sigma_B \sim \exp(\eta(B))$  and  $\sigma_A \perp \sigma_B$

Hence  $\mathbb{P}(\sigma_B < \sigma_{A \setminus B}) = \frac{\eta(B)}{\eta(B) + \eta(A \setminus B)} = \frac{\eta(B)}{\eta(A)}$

$\Rightarrow \epsilon_{\sigma_A} \sim \frac{\eta(\cdot; A)}{\eta(A)}$

