

# Meromorphic Lévy Processes and a new Wiener-Hopf Monte-Carlo simulation method

A. Kuznetsov, **A. E. Kyprianou**, J.C.Pardo and K. van Schaik

Department of Mathematical Sciences, University of Bath

## Motivation

## Motivation

- **Lévy process.** A (one dimensional) process with stationary and independent increments which has paths which are right continuous with left limits and therefore includes Brownian motion with drift, compound Poisson processes, stable processes amongst many others).

## Motivation

- **Lévy process.** A (one dimensional) process with stationary and independent increments which has paths which are right continuous with left limits and therefore includes Brownian motion with drift, compound Poisson processes, stable processes amongst many others).
- A popular model in mathematical finance for the evolution of a risky asset is

$$S_t := e^{X_t}, t \geq 0$$

where  $\{X_t : t \geq 0\}$  is a Lévy process.

## Motivation

- **Lévy process.** A (one dimensional) process with stationary and independent increments which has paths which are right continuous with left limits and therefore includes Brownian motion with drift, compound Poisson processes, stable processes amongst many others).
- A popular model in mathematical finance for the evolution of a risky asset is

$$S_t := e^{X_t}, t \geq 0$$

where  $\{X_t : t \geq 0\}$  is a Lévy process.

- Barrier options: The value of up-and-out barrier option with expiry date  $T$  and barrier  $b$  is typically priced as

$$\mathbb{E}_s(f(S_T)\mathbf{1}_{\{\bar{S}_T \leq b\}})$$

where  $\bar{S}_T = \sup_{u \leq T} S_u = \exp\{\sup_{u \leq T} X_u\}$ ,  $f$  is some nice function.

## Motivation

- **Lévy process.** A (one dimensional) process with stationary and independent increments which has paths which are right continuous with left limits and therefore includes Brownian motion with drift, compound Poisson processes, stable processes amongst many others).
- A popular model in mathematical finance for the evolution of a risky asset is

$$S_t := e^{X_t}, t \geq 0$$

where  $\{X_t : t \geq 0\}$  is a Lévy process.

- Barrier options: The value of up-and-out barrier option with expiry date  $T$  and barrier  $b$  is typically priced as

$$\mathbb{E}_s(f(S_T)\mathbf{1}_{\{\bar{S}_T \leq b\}})$$

where  $\bar{S}_T = \sup_{u \leq T} S_u = \exp\{\sup_{u \leq T} X_u\}$ ,  $f$  is some nice function.

- Other motivations from queuing theory, population models etc.

## Motivation

- **Lévy process.** A (one dimensional) process with stationary and independent increments which has paths which are right continuous with left limits and therefore includes Brownian motion with drift, compound Poisson processes, stable processes amongst many others).
- A popular model in mathematical finance for the evolution of a risky asset is

$$S_t := e^{X_t}, t \geq 0$$

where  $\{X_t : t \geq 0\}$  is a Lévy process.

- Barrier options: The value of up-and-out barrier option with expiry date  $T$  and barrier  $b$  is typically priced as

$$\mathbb{E}_s(f(S_T)\mathbf{1}_{\{\bar{S}_T \leq b\}})$$

where  $\bar{S}_T = \sup_{u \leq T} S_u = \exp\{\sup_{u \leq T} X_u\}$ ,  $f$  is some nice function.

- Other motivations from queuing theory, population models etc.
- One is fundamentally interested in the joint distribution

$$P(X_t \in dx, \bar{X}_t \in dy)$$

for any Lévy process  $(X, P)$ .

## Fourier methods



## Fourier methods

- Theoretically, things are already difficult enough if considering  $P(\overline{X}_t \in dy)$  or  $P(X_t \in dx)$ , especially in the former case.

## Fourier methods

- Theoretically, things are already difficult enough if considering  $P(\overline{X}_t \in dy)$  or  $P(X_t \in dx)$ , especially in the former case.
- For the case of  $P(X_t \in dx)$  one is sometimes lucky and knows this in explicit form. But usually one only knows something about

$$\begin{aligned}\Psi(\theta) &:= -\frac{1}{t} \log E(e^{i\theta X_t}) \\ &= ai\theta + \frac{1}{2}\sigma^2\theta^2 + \int_{\mathbb{R}} (1 - e^{i\theta x} + i\theta x \mathbf{1}_{\{|x| \leq 1\}}) \Pi(dx)\end{aligned}$$

where  $a \in \mathbb{R}$ ,  $\sigma \in \mathbb{R}$  and  $\Pi$  is a measure concentrated on  $\mathbb{R} \setminus \{0\}$  satisfying  $\int_{\mathbb{R}} (1 \wedge x^2) \Pi(dx) < \infty \dots$

## Fourier methods

- Theoretically, things are already difficult enough if considering  $P(\bar{X}_t \in dy)$  or  $P(X_t \in dx)$ , especially in the former case.
- For the case of  $P(X_t \in dx)$  one is sometimes lucky and knows this in explicit form. But usually one only knows something about

$$\begin{aligned}\Psi(\theta) &:= -\frac{1}{t} \log E(e^{i\theta X_t}) \\ &= ai\theta + \frac{1}{2}\sigma^2\theta^2 + \int_{\mathbb{R}} (1 - e^{i\theta x} + i\theta x \mathbf{1}_{\{|x| \leq 1\}}) \Pi(dx)\end{aligned}$$

where  $a \in \mathbb{R}$ ,  $\sigma \in \mathbb{R}$  and  $\Pi$  is a measure concentrated on  $\mathbb{R} \setminus \{0\}$  satisfying  $\int_{\mathbb{R}} (1 \wedge x^2) \Pi(dx) < \infty \dots$

- ...in which case there are fast-Fourier methods for inverting  $\exp\{-\Psi(\theta)t\}$  to give  $P(X_t \in dx)$ .

## Fourier methods

- Theoretically, things are already difficult enough if considering  $P(\overline{X}_t \in dy)$  or  $P(X_t \in dx)$ , especially in the former case.
- For the case of  $P(X_t \in dx)$  one is sometimes lucky and knows this in explicit form. But usually one only knows something about

$$\begin{aligned}\Psi(\theta) &:= -\frac{1}{t} \log E(e^{i\theta X_t}) \\ &= ai\theta + \frac{1}{2}\sigma^2\theta^2 + \int_{\mathbb{R}} (1 - e^{i\theta x} + i\theta x \mathbf{1}_{\{|x| \leq 1\}}) \Pi(dx)\end{aligned}$$

where  $a \in \mathbb{R}$ ,  $\sigma \in \mathbb{R}$  and  $\Pi$  is a measure concentrated on  $\mathbb{R} \setminus \{0\}$  satisfying  $\int_{\mathbb{R}} (1 \wedge x^2) \Pi(dx) < \infty \dots$

- ...in which case there are fast-Fourier methods for inverting  $\exp\{-\Psi(\theta)t\}$  to give  $P(X_t \in dx)$ .
- For the case of  $P(\overline{X}_t \in dx)$ , recent methods have concentrated on Fourier inversion of the, so called, Wiener-Hopf factors.

## Wiener-Hopf-Fourier methods

## Wiener-Hopf-Fourier methods

- Recall that it turns out that one may always uniquely decompose

$$\frac{q}{q + \Psi(\theta)} = E(e^{i\theta \bar{X}_{e_q}}) \times E(e^{i\theta \underline{X}_{e_q}})$$

where  $e_q$  is an independent and exponentially distributed random variable with rate  $q > 0$  and  $\underline{X}_t = \inf_{s \leq t} X_s$ .

## Wiener-Hopf-Fourier methods

- Recall that it turns out that one may always uniquely decompose

$$\frac{q}{q + \Psi(\theta)} = E(e^{i\theta \overline{X}_{e_q}}) \times E(e^{i\theta \underline{X}_{e_q}})$$

where  $e_q$  is an independent and exponentially distributed random variable with rate  $q > 0$  and  $\underline{X}_t = \inf_{s \leq t} X_s$ .

- *If one is in possession of close analytical expressions for these factors, Fourier inversion, first in  $\theta$  and then in  $q$  would be an option for accessing the law of  $\overline{X}_t$  and  $\underline{X}_t$ . However one is rarely in possession of the factors (even after 60 years of research into this topic), and even then there is the issue of the double Fourier inversion.*

## Wiener-Hopf-Fourier methods

- Recall that it turns out that one may always uniquely decompose

$$\frac{q}{q + \Psi(\theta)} = E(e^{i\theta \overline{X}_{e_q}}) \times E(e^{i\theta \underline{X}_{e_q}})$$

where  $e_q$  is an independent and exponentially distributed random variable with rate  $q > 0$  and  $\underline{X}_t = \inf_{s \leq t} X_s$ .

- *If one is in possession of close analytical expressions for these factors, Fourier inversion, first in  $\theta$  and then in  $q$  would be an option for accessing the law of  $\overline{X}_t$  and  $\underline{X}_t$ . However one is rarely in possession of the factors (even after 60 years of research into this topic), and even then there is the issue of the double Fourier inversion.*
- There are no convenient formulae which contain both  $X_t$  and  $\overline{X}_t$  which could be Fourier inverted.



## Some new ideas

## Some new ideas

- Suppose that  $e^{(1)}, e^{(2)}, \dots$  are a sequence of i.i.d unit mean exponentially distributed random variables.

## Some new ideas

- Suppose that  $e^{(1)}, e^{(2)}, \dots$  are a sequence of i.i.d unit mean exponentially distributed random variables.
- Note that

$$g(n, q) := \sum_{i=1}^n \frac{1}{q} e^{(i)}$$

is a Gamma (Erlang) distribution with parameters  $n$  and  $q$  and by the strong law of Large numbers, for  $t > 0$ ,

$$g(n, n/t) = \sum_{i=1}^n \frac{t}{n} e^{(i)} \rightarrow t$$

almost surely.

## Some new ideas

- Suppose that  $e^{(1)}, e^{(2)}, \dots$  are a sequence of i.i.d unit mean exponentially distributed random variables.
- Note that

$$\mathbf{g}(n, q) := \sum_{i=1}^n \frac{1}{q} e^{(i)}$$

is a Gamma (Erlang) distribution with parameters  $n$  and  $q$  and by the strong law of Large numbers, for  $t > 0$ ,

$$\mathbf{g}(n, n/t) = \sum_{i=1}^n \frac{t}{n} e^{(i)} \rightarrow t$$

almost surely.

- Hence for a suitably large  $n$ , we have in distribution

$$(X_{\mathbf{g}(n, n/t)}, \overline{X}_{\mathbf{g}(n, n/t)}) \simeq (X_t, \overline{X}_t).$$

Indeed since  $t$  is not a jump time with probability 1, we have that  $(X_{\mathbf{g}(n, n/t)}, \overline{X}_{\mathbf{g}(n, n/t)}) \rightarrow (X_t, \overline{X}_t)$  almost surely.

## Some new ideas

## Some new ideas

- A reformulation of the Wiener-Hopf factorization states that

$$X_{\mathbf{e}_q} \stackrel{d}{=} S_q + I_q$$

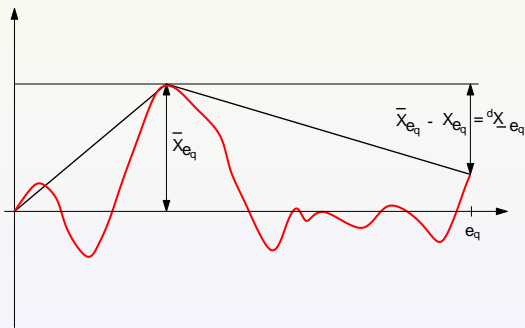
where  $S_q$  is independent of  $I_q$  and they are respectively equal in distribution to  $\overline{X}_{\mathbf{e}_q}$  and  $\underline{X}_{\mathbf{e}_q}$ .

## Some new ideas

- A reformulation of the Wiener-Hopf factorization states that

$$X_{e_q} \stackrel{d}{=} S_q + I_q$$

where  $S_q$  is independent of  $I_q$  and they are respectively equal in distribution to  $\bar{X}_{e_q}$  and  $\underline{X}_{e_q}$ .



## Some new ideas

- A reformulation of the Wiener-Hopf factorization states that

$$X_{\mathbf{e}_q} \stackrel{d}{=} S_q + I_q$$

where  $S_q$  is independent of  $I_q$  and they are respectively equal in distribution to  $\overline{X}_{\mathbf{e}_q}$  and  $\underline{X}_{\mathbf{e}_q}$ .

- Taking advantage of the above, the fact that  $X$  has stationary and independent increments and the fact that, as a time,  $\mathbf{g}(n, n/t)$  can be seen as  $n$  subsequent independent exponential time periods we have the following:



## Some new ideas

- **Theorem.** For all  $n \in \{1, 2, \dots\}$  and  $q > 0$ ,

$$(X_{\mathbf{g}(n,q)}, \bar{X}_{\mathbf{g}(n,q)}) \stackrel{d}{=} (V(n, q), J(n, q))$$

where

$$V(n, q) = \sum_{j=1}^n \{S_q^{(j)} + I_q^{(j)}\} \text{ and } J(n, q) := \bigvee_{i=0}^{n-1} \left( \sum_{j=1}^i \{S_q^{(j)} + I_q^{(j)}\} + S_q^{(i+1)} \right).$$

Here,  $S_q^{(0)} = I_q^{(0)} = 0$ ,  $\{S_q^{(j)} : j \geq 1\}$  are an i.i.d. sequence of random variables with common distribution equal to that of  $\bar{X}_{\mathbf{e}_q}$  and

$\{I_q^{(j)} : j \geq 1\}$  are another i.i.d. sequence of random variable with common distribution equal to that of  $\underline{X}_{\mathbf{e}_q}$ .

## Some new ideas

- Theorem.** For all  $n \in \{1, 2, \dots\}$  and  $q > 0$ ,

$$(X_{\mathbf{g}(n,q)}, \bar{X}_{\mathbf{g}(n,q)}) \stackrel{d}{=} (V(n, q), J(n, q))$$

where

$$V(n, q) = \sum_{j=1}^n \{S_q^{(j)} + I_q^{(j)}\} \text{ and } J(n, q) := \bigvee_{i=0}^{n-1} \left( \sum_{j=1}^i \{S_q^{(j)} + I_q^{(j)}\} + S_q^{(i+1)} \right).$$

Here,  $S_q^{(0)} = I_q^{(0)} = 0$ ,  $\{S_q^{(j)} : j \geq 1\}$  are an i.i.d. sequence of random variables with common distribution equal to that of  $\bar{X}_{\mathbf{e}_q}$  and

$\{I_q^{(j)} : j \geq 1\}$  are another i.i.d. sequence of random variable with common distribution equal to that of  $\underline{X}_{\mathbf{e}_q}$ .

- Moreover, we have the following obvious:

## Some new ideas

- **Theorem.** For all  $n \in \{1, 2, \dots\}$  and  $q > 0$ ,

$$(X_{\mathbf{g}(n,q)}, \overline{X}_{\mathbf{g}(n,q)}) \stackrel{d}{=} (V(n, q), J(n, q))$$

where

$$V(n, q) = \sum_{j=1}^n \{S_q^{(j)} + I_q^{(j)}\} \text{ and } J(n, q) := \bigvee_{i=0}^{n-1} \left( \sum_{j=1}^i \{S_q^{(j)} + I_q^{(j)}\} + S_q^{(i+1)} \right).$$

Here,  $S_q^{(0)} = I_q^{(0)} = 0$ ,  $\{S_q^{(j)} : j \geq 1\}$  are an i.i.d. sequence of random variables with common distribution equal to that of  $\overline{X}_{\mathbf{e}_q}$  and

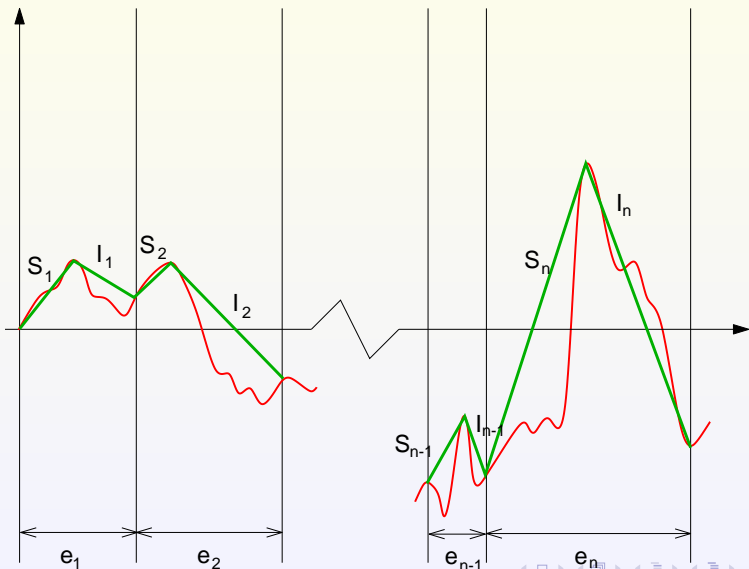
$\{I_q^{(j)} : j \geq 1\}$  are another i.i.d. sequence of random variable with common distribution equal to that of  $\underline{X}_{\mathbf{e}_q}$ .

- Moreover, we have the following obvious:
- **Corollary.** We have as  $n \uparrow \infty$

$$(V(n, n/t), J(n, n/t)) \rightarrow (X_t, \overline{X}_t)$$

where the convergence is understood in the distributional sense.

## Some new ideas



## Monte-Carlo simulation

## Monte-Carlo simulation

- The previous results suggest that to simulate, for example,  $\mathbb{E}(g(X_t, \bar{X}_t))$  one should follow the following algorithm:

## Monte-Carlo simulation

- The previous results suggest that to simulate, for example,  $\mathbb{E}(g(X_t, \bar{X}_t))$  one should follow the following algorithm:
- Sample independently from the distribution  $\bar{X}_{e_{n/t}}$  and  $\underline{X}_{e_{n/t}}$   $n \times m$ -times and then construct  $m$  independent versions of the variables  $V(n, n/t)$  and  $J(n, n/t)$ , say

$$\{V^{(i)}(n, n/t) : i = 1, \dots, m\} \text{ and } \{J^{(i)}(n, n/t) : i = 1, \dots, m\}.$$

## Monte-Carlo simulation

- The previous results suggest that to simulate, for example,  $\mathbb{E}(g(X_t, \bar{X}_t))$  one should follow the following algorithm:
- Sample independently from the distribution  $\bar{X}_{e_{n/t}}$  and  $\underline{X}_{e_{n/t}}$   $n \times m$ -times and then construct  $m$  independent versions of the variables  $V(n, n/t)$  and  $J(n, n/t)$ , say

$$\{V^{(i)}(n, n/t) : i = 1, \dots, m\} \text{ and } \{J^{(i)}(n, n/t) : i = 1, \dots, m\}.$$

- Then approximate

$$\mathbb{E}(g(X_t, \bar{X}_t)) \simeq \frac{1}{m} \sum_{i=1}^m g(V^{(i)}(n, n/t), J^{(i)}(n, n/t)).$$



## Monte-Carlo simulation

- The previous results suggest that to simulate, for example,  $\mathbb{E}(g(X_t, \bar{X}_t))$  one should follow the following algorithm:
- Sample independently from the distribution  $\bar{X}_{e_{n/t}}$  and  $\underline{X}_{e_{n/t}}$   $n \times m$ -times and then construct  $m$  independent versions of the variables  $V(n, n/t)$  and  $J(n, n/t)$ , say

$$\{V^{(i)}(n, n/t) : i = 1, \dots, m\} \text{ and } \{J^{(i)}(n, n/t) : i = 1, \dots, m\}.$$

- Then approximate

$$\mathbb{E}(g(X_t, \bar{X}_t)) \simeq \frac{1}{m} \sum_{i=1}^m g(V^{(i)}(n, n/t), J^{(i)}(n, n/t)).$$

- This numerical procedure has disposed of one (numerical) Fourier inverse computation.

## Monte-Carlo simulation

- The previous results suggest that to simulate, for example,  $\mathbb{E}(g(X_t, \bar{X}_t))$  one should follow the following algorithm:
- Sample independently from the distribution  $\bar{X}_{e_{n/t}}$  and  $\underline{X}_{e_{n/t}}$   $n \times m$ -times and then construct  $m$  independent versions of the variables  $V(n, n/t)$  and  $J(n, n/t)$ , say

$$\{V^{(i)}(n, n/t) : i = 1, \dots, m\} \text{ and } \{J^{(i)}(n, n/t) : i = 1, \dots, m\}.$$

- Then approximate

$$\mathbb{E}(g(X_t, \bar{X}_t)) \simeq \frac{1}{m} \sum_{i=1}^m g(V^{(i)}(n, n/t), J^{(i)}(n, n/t)).$$

- This numerical procedure has disposed of one (numerical) Fourier inverse computation.
- This still leaves the problem of simulating from the unknown distribution  $\bar{X}_{e_{n/t}}$  and  $\underline{X}_{e_{n/t}}$  i.e. we are still one (numerical) Fourier transform away from  $(X_t, \bar{X}_t)$

## Meromorphic Lévy processes (definition)

## Meromorphic Lévy processes (definition)

- A Lévy process is said to belong to the Meromorphic class ( $M$ -class), if and only if the Lévy measure  $\Pi(dx)$  has a density with respect to the Lebesgue measure, given by

$$\pi(x) = \mathbb{I}_{\{x>0\}} \sum_{n \geq 1} a_n \rho_n e^{-\rho_n x} + \mathbb{I}_{\{x<0\}} \sum_{n \geq 1} \hat{a}_n \hat{\rho}_n e^{\hat{\rho}_n x}, \quad (1)$$

where all the coefficients  $a_n, \hat{a}_n, \rho_n, \hat{\rho}_n$  are positive, the sequences  $\{\rho_n\}_{n \geq 1}$  and  $\{\hat{\rho}_n\}_{n \geq 1}$  are strictly increasing, and  $\rho_n \rightarrow +\infty$  and  $\hat{\rho}_n \rightarrow +\infty$  as  $n \rightarrow +\infty$ .

## Meromorphic Lévy processes (definition)

- A Lévy process is said to belong to the Meromorphic class ( $M$ -class), if and only if the Lévy measure  $\Pi(dx)$  has a density with respect to the Lebesgue measure, given by

$$\pi(x) = \mathbb{I}_{\{x>0\}} \sum_{n \geq 1} a_n \rho_n e^{-\rho_n x} + \mathbb{I}_{\{x<0\}} \sum_{n \geq 1} \hat{a}_n \hat{\rho}_n e^{\hat{\rho}_n x}, \quad (1)$$

where all the coefficients  $a_n, \hat{a}_n, \rho_n, \hat{\rho}_n$  are positive, the sequences  $\{\rho_n\}_{n \geq 1}$  and  $\{\hat{\rho}_n\}_{n \geq 1}$  are strictly increasing, and  $\rho_n \rightarrow +\infty$  and  $\hat{\rho}_n \rightarrow +\infty$  as  $n \rightarrow +\infty$ .

- We allow the case of a finite number summands (on either or both sides of the origin) with obvious modifications to the above.

## Meromorphic Lévy processes (definition)

- A Lévy process is said to belong to the Meromorphic class ( $M$ -class), if and only if the Lévy measure  $\Pi(dx)$  has a density with respect to the Lebesgue measure, given by

$$\pi(x) = \mathbb{I}_{\{x>0\}} \sum_{n \geq 1} a_n \rho_n e^{-\rho_n x} + \mathbb{I}_{\{x<0\}} \sum_{n \geq 1} \hat{a}_n \hat{\rho}_n e^{\hat{\rho}_n x}, \quad (1)$$

where all the coefficients  $a_n, \hat{a}_n, \rho_n, \hat{\rho}_n$  are positive, the sequences  $\{\rho_n\}_{n \geq 1}$  and  $\{\hat{\rho}_n\}_{n \geq 1}$  are strictly increasing, and  $\rho_n \rightarrow +\infty$  and  $\hat{\rho}_n \rightarrow +\infty$  as  $n \rightarrow +\infty$ .

- We allow the case of a finite number summands (on either or both sides of the origin) with obvious modifications to the above.
- To ensure that  $\int_{\mathbb{R}} x^2 \pi(x) dx$  converges we need to impose the additional constraint that

$$\sum_{n \geq 1} a_n \rho_n^{-2} + \sum_{n \geq 1} \hat{a}_n \hat{\rho}_n^{-2} < \infty$$

## Meromorphic Lévy processes (equivalent definition)

- (i) The characteristic exponent  $\Psi(z)$  is a meromorphic function which has poles at points  $\{-i\rho_n, i\hat{\rho}_n\}_{n \geq 1}$ , where  $\rho_n$  and  $\hat{\rho}_n$  are positive real numbers.
- (ii) For  $q \geq 0$  function  $q + \Psi(z)$  has roots at points  $\{-i\zeta_n, i\hat{\zeta}_n\}_{n \geq 1}$  where  $\zeta_n$  and  $\hat{\zeta}_n$  are nonnegative real numbers (strictly positive if  $q > 0$ ). We will write  $\zeta_n(q)$ ,  $\hat{\zeta}_n(q)$  if we need to stress the dependence on  $q$ .
- (iii) The roots and poles of  $q + \Psi(iz)$  satisfy the following interlacing condition

$$\dots - \rho_2 < -\zeta_2 < -\rho_1 < -\zeta_1 < 0 < \hat{\zeta}_1 < \hat{\rho}_1 < \hat{\zeta}_2 < \hat{\rho}_2 < \dots$$

- (iv) There exists  $\alpha > 1/2$  such that  $\rho_n \sim cn^\alpha$  and  $\hat{\rho}_n \sim \hat{c}n^\alpha$  as  $n \rightarrow +\infty$ .
- (v) The Wiener-Hopf factors are expressed as convergent infinite products,

$$\mathbb{E} \left[ e^{-z\bar{X}_{\mathbf{e}_q}} \right] = \prod_{n \geq 1} \frac{1 + \frac{z}{\rho_n}}{1 + \frac{z}{\zeta_n}}$$

$$\mathbb{E} \left[ e^{zX_{\mathbf{e}_q}} \right] = \prod_{n \geq 1} \frac{1 + \frac{z}{\hat{\rho}_n}}{1 + \frac{z}{\hat{\zeta}_n}}.$$

## Example: hyper-exponential jumps



## Example: hyper-exponential jumps

- The density of the Lévy measure is

$$\pi(x) = \mathbf{1}_{\{x>0\}} \sum_{i=1}^N a_i \rho_i e^{-\rho_i x} + \mathbf{1}_{\{x<0\}} \sum_{i=1}^{\hat{N}} \hat{a}_i \hat{\rho}_i e^{\hat{\rho}_i x},$$

where  $a_i$ ,  $\hat{a}_i$ ,  $\rho_i$  and  $\hat{\rho}_i$  are positive numbers.

## Example: hyper-exponential jumps

- The density of the Lévy measure is

$$\pi(x) = \mathbf{1}_{\{x>0\}} \sum_{i=1}^N a_i \rho_i e^{-\rho_i x} + \mathbf{1}_{\{x<0\}} \sum_{i=1}^{\hat{N}} \hat{a}_i \hat{\rho}_i e^{\hat{\rho}_i x},$$

where  $a_i$ ,  $\hat{a}_i$ ,  $\rho_i$  and  $\hat{\rho}_i$  are positive numbers.

- Including Gaussian and linear drift, one can verify that the characteristic exponent is a rational function and that hyper-exponential Lévy processes have finite activity jumps and paths of bounded variation unless  $\sigma > 0$ .

## Example: hyper-exponential jumps

- The density of the Lévy measure is

$$\pi(x) = \mathbf{1}_{\{x>0\}} \sum_{i=1}^N a_i \rho_i e^{-\rho_i x} + \mathbf{1}_{\{x<0\}} \sum_{i=1}^{\hat{N}} \hat{a}_i \hat{\rho}_i e^{\hat{\rho}_i x},$$

where  $a_i$ ,  $\hat{a}_i$ ,  $\rho_i$  and  $\hat{\rho}_i$  are positive numbers.

- Including Gaussian and linear drift, one can verify that the characteristic exponent is a rational function and that hyper-exponential Lévy processes have finite activity jumps and paths of bounded variation unless  $\sigma > 0$ .
- Note that this class has been looked at by many other authors in the past and historically it starts life as the Kou process.

## Example: Kuznetsov's $\beta$ -family

## Example: Kuznetsov's $\beta$ -family

- The characteristic exponent ( $\Psi(\theta) = -\log \mathbb{E}(e^{i\theta X_1}), \theta \in \mathbb{R}$ ) is given by

$$\begin{aligned} \Psi(\theta) = & \ iaz + \frac{1}{2}\sigma^2 z^2 + \frac{c_1}{\beta_1} \left\{ \mathbf{B}(\alpha_1, 1 - \lambda_1) - \mathbf{B}\left(\alpha_1 - \frac{i\theta}{\beta_1}, 1 - \lambda_1\right) \right\} \\ & + \frac{c_2}{\beta_2} \left\{ \mathbf{B}(\alpha_2, 1 - \lambda_2) - \mathbf{B}\left(\alpha_2 + \frac{i\theta}{\beta_2}, 1 - \lambda_2\right) \right\} \end{aligned}$$

where  $\mathbf{B}(x, y) = \Gamma(x)\Gamma(y)/\Gamma(x + y)$  is the Beta function, with parameter range  $a \in \mathbb{R}, \sigma, c_i, \alpha_i, \beta_i > 0$  and  $\lambda_1, \lambda_2 \in (0, 3) \setminus \{1, 2\}$ .

## Example: Kuznetsov's $\beta$ -family

- The characteristic exponent ( $\Psi(\theta) = -\log \mathbb{E}(e^{i\theta X_1}), \theta \in \mathbb{R}$ ) is given by

$$\begin{aligned} \Psi(\theta) = & \ iaz + \frac{1}{2}\sigma^2 z^2 + \frac{c_1}{\beta_1} \left\{ \mathbf{B}(\alpha_1, 1 - \lambda_1) - \mathbf{B}\left(\alpha_1 - \frac{i\theta}{\beta_1}, 1 - \lambda_1\right) \right\} \\ & + \frac{c_2}{\beta_2} \left\{ \mathbf{B}(\alpha_2, 1 - \lambda_2) - \mathbf{B}\left(\alpha_2 + \frac{i\theta}{\beta_2}, 1 - \lambda_2\right) \right\} \end{aligned}$$

where  $\mathbf{B}(x, y) = \Gamma(x)\Gamma(y)/\Gamma(x+y)$  is the Beta function, with parameter range  $a \in \mathbb{R}, \sigma, c_i, \alpha_i, \beta_i > 0$  and  $\lambda_1, \lambda_2 \in (0, 3) \setminus \{1, 2\}$ .

- The corresponding Lévy measure  $\Pi$  has density

$$\pi(x) = c_1 \frac{e^{-\alpha_1 \beta_1 x}}{(1 - e^{-\beta_1 x})^{\lambda_1}} \mathbf{1}_{\{x > 0\}} + c_2 \frac{e^{\alpha_2 \beta_2 x}}{(1 - e^{\beta_2 x})^{\lambda_2}} \mathbf{1}_{\{x < 0\}}.$$

The  $\beta$ -class of Lévy processes includes another recently introduced family of Lévy processes known as Lamperti-stable processes.

## Example: Hypergeometric Lévy processes

## Example: Hypergeometric Lévy processes

- The characteristic exponent ( $\Psi(\theta) = \mathbb{E}(e^{i\theta X_1}), \theta \in \mathbb{R}$ ) is given by

$$\Psi(\theta) = \frac{\Gamma(1 - \beta + \gamma - i\theta)}{\Gamma(1 - \beta + i\theta)} \frac{\Gamma(\hat{\beta} + \hat{\gamma} + i\theta)}{\Gamma(\hat{\beta} + i\theta)}$$

where  $(\beta, \gamma, \hat{\beta}, \hat{\gamma})$  belong to the admissible range

$$\{\beta \leq 1, \gamma \in (0, 1), \hat{\beta} \geq 0, \hat{\gamma} \in (0, 1)\}.$$



## Example: Hypergeometric Lévy processes

- The characteristic exponent ( $\Psi(\theta) = \mathbb{E}(e^{i\theta X_1}), \theta \in \mathbb{R}$ ) is given by

$$\Psi(\theta) = \frac{\Gamma(1 - \beta + \gamma - i\theta)}{\Gamma(1 - \beta + i\theta)} \frac{\Gamma(\hat{\beta} + \hat{\gamma} + i\theta)}{\Gamma(\hat{\beta} + i\theta)}$$

where  $(\beta, \gamma, \hat{\beta}, \hat{\gamma})$  belong to the admissible range

$$\{\beta \leq 1, \gamma \in (0, 1), \hat{\beta} \geq 0, \hat{\gamma} \in (0, 1)\}.$$

- The Lévy density is given by

$$\pi(x) = \begin{cases} -\frac{\Gamma(\eta)}{\Gamma(\eta - \hat{\gamma})\Gamma(-\gamma)} e^{-(1-\beta+\gamma)x} {}_2F_1(1 + \gamma, \eta; \eta - \hat{\gamma}; e^{-x}) & \text{if } x > 0 \\ -\frac{\Gamma(\eta)}{\Gamma(\eta - \gamma)\Gamma(-\hat{\gamma})} e^{(\hat{\beta} + \hat{\gamma})x} {}_2F_1(1 + \hat{\gamma}, \eta; \eta - \gamma; e^{-x}) & \text{if } x < 0 \end{cases}$$

where  $\eta = 1 - \beta + \gamma + \hat{\beta} + \hat{\gamma}$ .

## Distribution of extrema

## Distribution of extrema

- For  $x \geq 0$

$$\begin{aligned}\mathbb{P}(\overline{X}_{e_q} \in dx) &= \bar{a}(\rho, \zeta)^T \times \bar{v}(\zeta, x) dx \\ \mathbb{P}(-\underline{X}_{e_q} \in dx) &= \bar{a}(\hat{\rho}, \hat{\zeta})^T \times \bar{v}(\hat{\zeta}, x) dx.\end{aligned}$$

## Distribution of extrema

- For  $x \geq 0$

$$\begin{aligned}\mathbb{P}(\overline{X}_{e_q} \in dx) &= \bar{a}(\rho, \zeta)^T \times \bar{v}(\zeta, x) dx \\ \mathbb{P}(-\underline{X}_{e_q} \in dx) &= \bar{a}(\hat{\rho}, \hat{\zeta})^T \times \bar{v}(\hat{\zeta}, x) dx.\end{aligned}$$

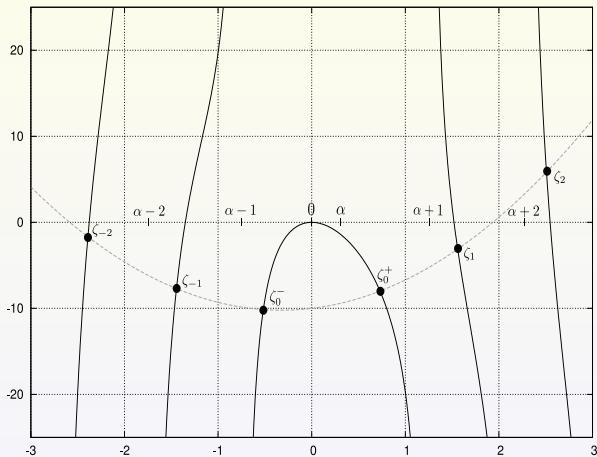
- Here  $\bar{a}(\rho, \zeta) = [a_0(\rho, \zeta), a_1(\rho, \zeta), a_2(\rho, \zeta), \dots]^T$  such that

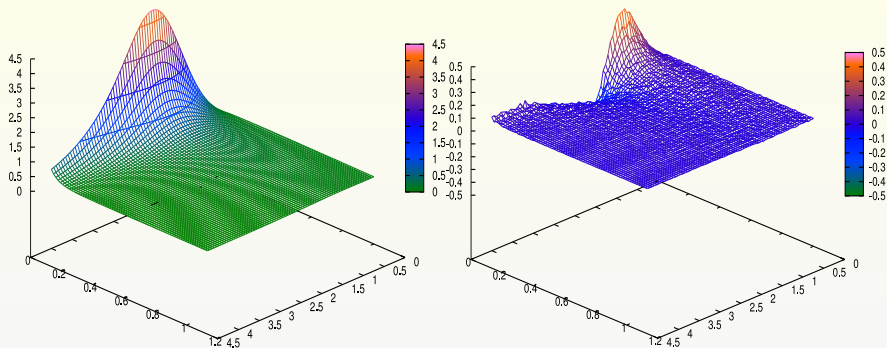
$$a_0(\rho, \zeta) = \lim_{n \rightarrow +\infty} \prod_{k=1}^n \frac{\zeta_k}{\rho_k}, \quad a_n(\rho, \zeta) = \left(1 - \frac{\zeta_n}{\rho_n}\right) \prod_{\substack{k \geq 1 \\ k \neq n}} \frac{1 - \frac{\zeta_n}{\rho_k}}{1 - \frac{\zeta_n}{\zeta_k}}$$

$$\bar{v}(\zeta, x) = \left[ \delta_0(x), \zeta_1 e^{-\zeta_1 x}, \zeta_2 e^{-\zeta_2 x}, \dots \right]^T,$$

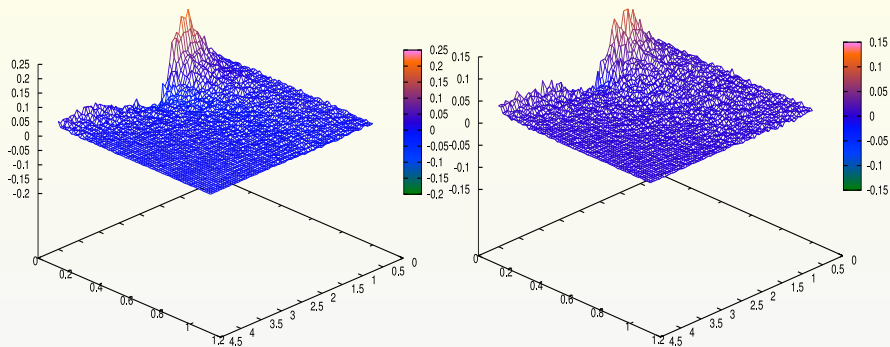
where  $\delta_0(x)$  is the Dirac delta function at  $x = 0$ . A similar expression holds for  $\bar{a}(\rho, \zeta)$ .

## Computing roots

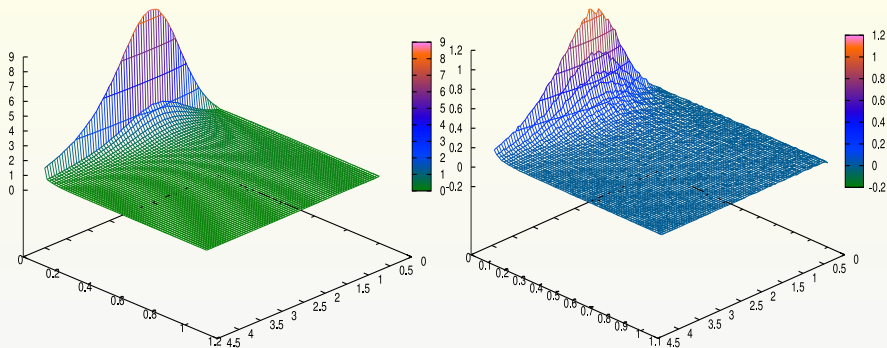




**Figure:** Computing the joint density of  $(\bar{X}_1, \bar{X}_1 - X_1)$  for parameter Set 1. Here  $\bar{X}_1 \in [0, 1]$  and  $\bar{X}_1 - X_1 \in [0, 4]$ . Fourier method benchmark (left),  $N = 20$ , WH-MC error (right).



**Figure:** Computing the joint density of  $(\bar{X}_1, \bar{X}_1 - X_1)$  for parameter Set 1. Here  $\bar{X}_1 \in [0, 1]$  and  $\bar{X}_1 - X_1 \in [0, 4]$ .  $N = 50$ , WH-MC error(left),  $N = 100$ , WH-MC error (right).



**Figure:** Computing the joint density of  $(\bar{X}_1, \bar{X}_1 - X_1)$  for parameter Set 1. Here  $\bar{X}_1 \in [0, 1]$  and  $\bar{X}_1 - X_1 \in [0, 4]$ .  $N = 100$ , MC simulation (left),  $N = 100$ , MC error (right).



## Advantages of the WH-MC method over RW approximation MC

## Advantages of the WH-MC method over RW approximation MC

- Total computation time for WH-MC is at most half the computation time for Fourier inversion of  $\exp -\Psi(z)$  followed by a random walk simulation.

## Advantages of the WH-MC method over RW approximation MC

- Total computation time for WH-MC is at most half the computation time for Fourier inversion of  $\exp -\Psi(z)$  followed by a random walk simulation.
- The overwhelming majority of the WH-MC method is the simulation, computing the roots takes 1% of the time. Roots can be stored once they have been computed.

## Advantages of the WH-MC method over RW approximation MC

- Total computation time for WH-MC is at most half the computation time for Fourier inversion of  $\exp -\Psi(z)$  followed by a random walk simulation.
- The overwhelming majority of the WH-MC method is the simulation, computing the roots takes 1% of the time. Roots can be stored once they have been computed.
- Considerably more accurate for the same number of steps in each cycle.

## Advantages of the WH-MC method over RW approximation MC

- Total computation time for WH-MC is at most half the computation time for Fourier inversion of  $\exp -\Psi(z)$  followed by a random walk simulation.
- The overwhelming majority of the WH-MC method is the simulation, computing the roots takes 1% of the time. Roots can be stored once they have been computed.
- Considerably more accurate for the same number of steps in each cycle.
- Does not artificially build in an atom at zero in the numerical distribution of  $\bar{X}_t$ .

## More identities: One sided exit problem

## More identities: One sided exit problem

- Define a matrix  $\mathbf{A} = \{a_{i,j}\}_{i,j \geq 0}$  as

$$a_{i,j} = \begin{cases} 0 & \text{if } i = 0, j \geq 0 \\ \mathbf{a}_i(\rho, \zeta) \mathbf{b}_0(\zeta, \rho) & \text{if } i \geq 1, j = 0 \\ \frac{\mathbf{a}_i(\rho, \zeta) \mathbf{b}_j(\zeta, \rho)}{\rho_j - \zeta_i} & \text{if } i \geq 1, j \geq 1 \end{cases} \quad (2)$$

Then for  $c > 0$  and  $y \geq 0$  we have

$$\mathbb{E} \left[ e^{-q\tau_c^+} \mathbb{I} \left( X_{\tau_c^+} - c \in dy \right) \right] = \bar{\mathbf{v}}(\zeta, c)^T \times \mathbf{A} \times \bar{\mathbf{v}}(\rho, y) dy. \quad (3)$$

## More identities: Two-sided exit problem



## More identities: Two-sided exit problem

- Let  $a > 0$  and define a matrix  $\mathbf{B} = \mathbf{B}(\hat{\rho}, \zeta, a) = \{b_{i,j}\}_{i,j \geq 0}$  with

$$b_{i,j} = \begin{cases} \zeta_j e^{-a\zeta_j} & \text{if } i = 0, j \geq 1 \\ 0 & \text{if } i \geq 0, j = 0 \\ \frac{\hat{\rho}_i \zeta_j}{\hat{\rho}_i + \zeta_j} e^{-a\zeta_j} & \text{if } i \geq 1, j \geq 1 \end{cases}$$

and similarly  $\hat{\mathbf{B}} = \mathbf{B}(\rho, \hat{\zeta}, a)$ . There exist matrices  $\mathbf{C}_1, \mathbf{C}_2$  such that for  $x \in (0, a)$  we have

$$\begin{aligned} & \mathbb{E}_x \left[ e^{-q\tau_a^+} \mathbb{I} \left( X_{\tau_a^+} \in dy ; \tau_a^+ < \tau_0^- \right) \right] \\ &= \left[ \bar{\mathbf{v}}(\zeta, a-x)^T \times \mathbf{C}_1 + \bar{\mathbf{v}}(\hat{\zeta}, x)^T \times \mathbf{C}_2 \right] \times \bar{\mathbf{v}}(\rho, y-a) dy \end{aligned}$$

## More identities: Two-sided exit problem

- Let  $a > 0$  and define a matrix  $\mathbf{B} = \mathbf{B}(\hat{\rho}, \zeta, a) = \{b_{i,j}\}_{i,j \geq 0}$  with

$$b_{i,j} = \begin{cases} \zeta_j e^{-a\zeta_j} & \text{if } i = 0, j \geq 1 \\ 0 & \text{if } i \geq 0, j = 0 \\ \frac{\hat{\rho}_i \zeta_j}{\hat{\rho}_i + \zeta_j} e^{-a\zeta_j} & \text{if } i \geq 1, j \geq 1 \end{cases}$$

and similarly  $\hat{\mathbf{B}} = \mathbf{B}(\rho, \hat{\zeta}, a)$ . There exist matrices  $\mathbf{C}_1, \mathbf{C}_2$  such that for  $x \in (0, a)$  we have

$$\begin{aligned} \mathbb{E}_x \left[ e^{-q\tau_a^+} \mathbb{I} \left( X_{\tau_a^+} \in dy ; \tau_a^+ < \tau_0^- \right) \right] \\ = \left[ \bar{v}(\zeta, a-x)^T \times \mathbf{C}_1 + \bar{v}(\hat{\zeta}, x)^T \times \mathbf{C}_2 \right] \times \bar{v}(\rho, y-a) dy \end{aligned}$$

- These matrices satisfy the following system of linear equations

$$\begin{cases} \mathbf{C}_1 &= \mathbf{A} - \hat{\mathbf{C}}_2 \mathbf{B} \mathbf{A} \\ \hat{\mathbf{C}}_2 &= -\mathbf{C}_1 \hat{\mathbf{B}} \mathbf{A} \end{cases} \quad \begin{cases} \hat{\mathbf{C}}_1 &= \hat{\mathbf{A}} - \mathbf{C}_2 \hat{\mathbf{B}} \hat{\mathbf{A}} \\ \mathbf{C}_2 &= -\hat{\mathbf{C}}_1 \mathbf{B} \hat{\mathbf{A}} \end{cases}$$

This system of linear equations can be solved iteratively with exponential convergence.

## More identities: Half-line resolvent

## More identities: Half-line resolvent

- For  $a > 0$ ,  $y \leq a$  we define  $R^{(q)}(a, dy) := \int_0^\infty e^{-qt} \mathbb{P}(X_t \in dy; t < \tau_a^+) dt$

## More identities: Half-line resolvent

- For  $a > 0$ ,  $y \leq a$  we define  $R^{(q)}(a, dy) := \int_0^\infty e^{-qt} \mathbb{P}(X_t \in dy; t < \tau_a^+) dt$
- Define a matrix  $\mathbf{D} = \{d_{i,j}\}_{i,j \geq 0}$  as follows

$$d_{i,j} = \begin{cases} 0 & \text{if } i = 0 \text{ or } j = 0 \\ \mathbf{a}_i(\rho, \zeta) \frac{\zeta_i \hat{\zeta}_j}{\zeta_i + \hat{\zeta}_j} \mathbf{a}_j(\hat{\rho}, \hat{\zeta}) & \text{if } i \geq 1, j \geq 1 \end{cases}$$

## More identities: Half-line resolvent

- For  $a > 0$ ,  $y \leq a$  we define  $R^{(q)}(a, dy) := \int_0^\infty e^{-qt} \mathbb{P}(X_t \in dy; t < \tau_a^+) dt$
- Define a matrix  $\mathbf{D} = \{d_{i,j}\}_{i,j \geq 0}$  as follows

$$d_{i,j} = \begin{cases} 0 & \text{if } i = 0 \text{ or } j = 0 \\ \mathbf{a}_i(\rho, \zeta) \frac{\zeta_i \hat{\zeta}_j}{\zeta_i + \hat{\zeta}_j} \mathbf{a}_j(\hat{\rho}, \hat{\zeta}) & \text{if } i \geq 1, j \geq 1 \end{cases}$$

- Then if  $y \leq a$  we have

$$qR^{(q)}(a, dy) = [\bar{\mathbf{v}}(\zeta, 0 \vee y) \times \mathbf{D} \times \bar{\mathbf{v}}(\hat{\zeta}, 0 \vee (-y))] - \bar{\mathbf{v}}(\zeta, a) \times \mathbf{D} \times \bar{\mathbf{v}}(\hat{\zeta}, a - y)] dy.$$

## Example of numerics

## Example of numerics

- Choose an example from Kuznetsov's  $\beta$ -class that has bounded variation jump component and concentrate on four cases: With/without Gaussian component, drift to  $\pm\infty$ .



## Example of numerics

- Choose an example from Kuznetsov's  $\beta$ -class that has bounded variation jump component and concentrate on four cases: With/without Gaussian component, drift to  $\pm\infty$ .
- For the above four cases, consider the following densities.

- (i) density of the overshoot if the exit happens at the upper boundary

$$f_1(x, y) = \frac{d}{dy} \mathbb{E}_x \left[ e^{-q\tau_1^+} \mathbb{I} \left( X_{\tau_1^+} \leq y ; \tau_1^+ < \tau_0^- \right) \right]$$

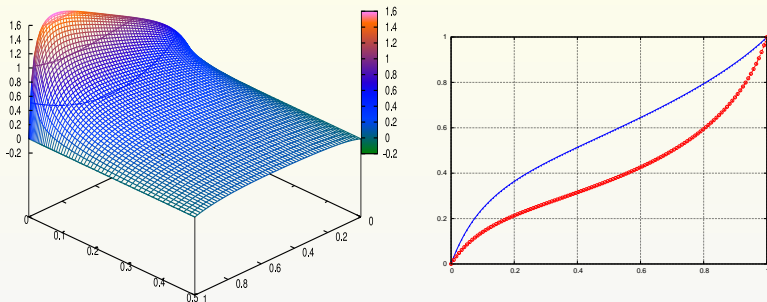
- (ii) probability of exiting from the interval  $[0, 1]$  at the upper boundary

$$f_2(x) = \mathbb{E}_x \left[ e^{-q\tau_1^+} \mathbb{I} \left( \tau_1^+ < \tau_0^- \right) \right]$$

- (iii) probability of exiting the interval  $[0, 1]$  by creeping across the upper boundary

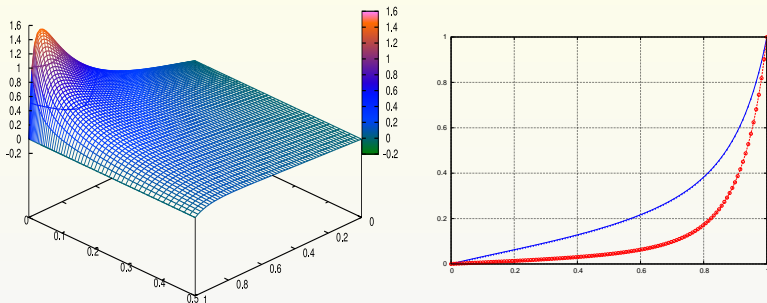
$$f_3(x) = \mathbb{E}_x \left[ e^{-q\tau_1^+} \mathbb{I} \left( X_{\tau_1^+} = 1 ; \tau_1^+ < \tau_0^- \right) \right]$$

## Double sided exit: $\sigma > 0$ and positive drift



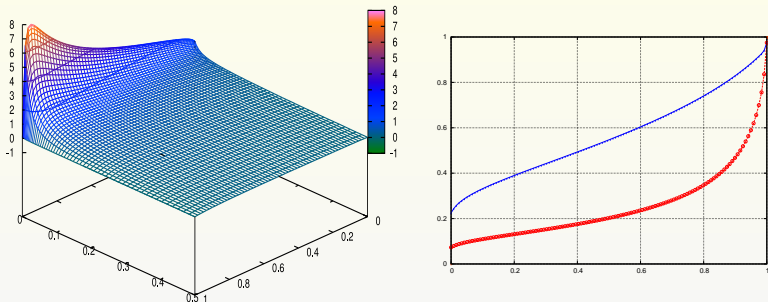
**Figure:** Unbounded variation case ( $\sigma = 0.5$ ): computing the density of the overshoot  $f_1(x, y)$  ( $x \in (0, 1)$ ,  $y \in (0, 0.5)$ ), probability of first exit  $f_2(x)$  and probability of creeping  $f_3(x)$  for parameter Set 1, positive drift  $\mu = 1$

## Double sided exit: $\sigma > 0$ and negative drift



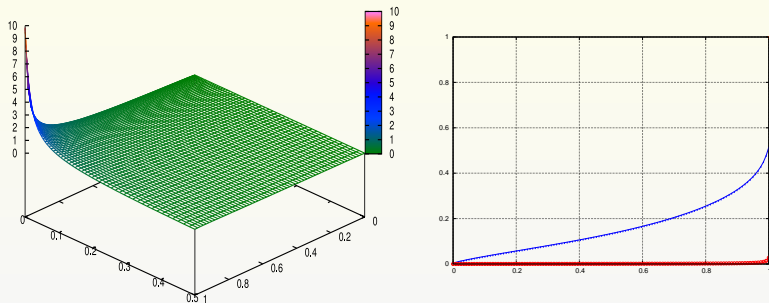
**Figure:** Unbounded variation case ( $\sigma = 0.5$ ): computing the density of the overshoot  $f_1(x, y)$  ( $x \in (0, 1)$ ,  $y \in (0, 0.5)$ ), probability of first exit  $f_2(x)$  and probability of creeping  $f_3(x)$  for parameter Set 2, negative drift  $\mu = -1$ .

## Double sided exit: bounded variation and positive drift



**Figure:**  $\sigma = 0$ , positive drift: computing the density of the overshoot  $f_1(x, y)$  ( $x \in (0, 1)$ ,  $y \in (0, 0.5)$ ), probability of first exit  $f_2(x)$  and probability of creeping  $f_3(x)$ .

## Double sided exit: bounded variation and negative drift



**Figure:**  $\sigma = 0$ , negative drift: computing the density of the overshoot  $f_1(x, y)$  ( $x \in (0, 1)$ ,  $y \in (0, 0.5)$ ), probability of first exit  $f_2(x)$  and probability of creeping  $f_3(x)$ .

## Simulating processes with heavy tails

## Simulating processes with heavy tails

- A little thought shows that a huge class of Lévy processes can be written as the independent sum of a  $\beta$ -process plus and independent compound Poisson process. Say,

$$Y_t = X_t + \sum_{i=1}^{N_t} \xi_i$$

where  $\{N_t : t \geq 0\}$  is a Poisson process of rate  $\gamma$  and  $\{\xi_i : i \geq 1\}$  and i.i.d. sequence.

## Simulating processes with heavy tails

- A little thought shows that a huge class of Lévy processes can be written as the independent sum of a  $\beta$ -process plus and independent compound Poisson process. Say,

$$Y_t = X_t + \sum_{i=1}^{N_t} \xi_i$$

where  $\{N_t : t \geq 0\}$  is a Poisson process of rate  $\gamma$  and  $\{\xi_i : i \geq 1\}$  and i.i.d. sequence.

- Define iteratively for  $n \geq 1$

$$V(n, \lambda) = V(n-1, \lambda) + S_{\lambda+\gamma}^{(n)} + I_{\lambda+\gamma}^{(n)} + \xi_n(1 - \beta_n)$$

$$J(n, \lambda) = \max \left( V(n, \lambda), J(n-1, \lambda), V(n-1, \lambda) + S_{\lambda+\gamma}^{(n)} \right)$$

where  $V(0, \lambda) = J(0, \lambda) = 0$  and  $\{\beta_n : n \geq 1\}$  are an i.i.d. sequence of Bernoulli random variables such that  $\mathbb{P}(\beta_n = 1) = \lambda/(\gamma + \lambda)$ . Then

$$(Y_{\mathbf{g}(n,\lambda)}, \bar{Y}_{\mathbf{g}(n,\lambda)}) \stackrel{d}{=} (V(T_n, \lambda), J(T_n, \lambda))$$

where  $T_n = \min\{j \geq 1 : \sum_{i=1}^j \beta_i = n\}$ .



## Approximate simulation of the law of $(X_t, \bar{X}_t, \underline{X}_t)$

Define iteratively for  $n \geq 1$

$$V(n, \lambda) = V(n-1, \lambda) + S_\lambda^{(n)} + I_\lambda^{(n)}$$

$$J(n, \lambda) = \max\left(J(n-1, \lambda), V(n-1, \lambda) + S_\lambda^{(n)}\right)$$

$$K(n, \lambda) = \min(K(n-1, \lambda), V(n, \lambda))$$

$$\tilde{J}(n, \lambda) = \max(\tilde{J}(n-1, \lambda), V(n, \lambda))$$

$$\tilde{K}(n, \lambda) = \min\left(\tilde{K}(n-1, \lambda), V(n-1, \lambda) + I_\lambda^{(n)}\right),$$

where  $V(0, \lambda) = J(0, \lambda) = K(0, \lambda) = \tilde{J}(0, \lambda) = \tilde{K}(0, \lambda) = 0$ . Then for any bounded function  $f(x, y, z) : \mathbb{R}^3 \mapsto \mathbb{R}$  which is increasing in  $z$ -variable we have

$$\mathbb{E}[f(V(n, \lambda), J(n, \lambda), K(n, \lambda))] \geq \mathbb{E}[f(X_{\mathbf{g}(n, \lambda)}, \bar{X}_{\mathbf{g}(n, \lambda)}, \underline{X}_{\mathbf{g}(n, \lambda)})]$$

$$\mathbb{E}[f(V(n, \lambda), \tilde{K}(n, \lambda), \tilde{J}(n, \lambda))] \leq \mathbb{E}[f(X_{\mathbf{g}(n, \lambda)}, \underline{X}_{\mathbf{g}(n, \lambda)}, \bar{X}_{\mathbf{g}(n, \lambda)})].$$