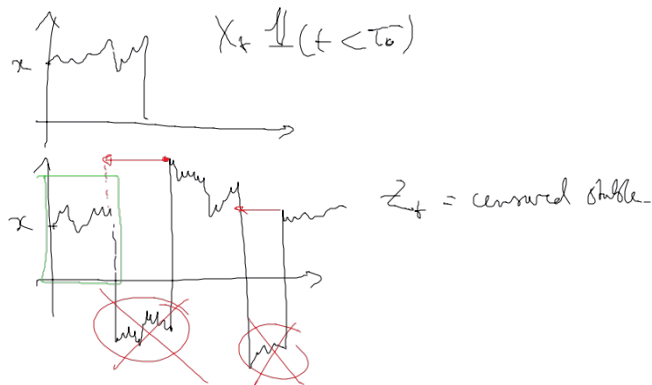


Censored Stable



Where $X_t \mathbb{1}(t < t_0)$ kills a L.p. (in the Lamperti transform) Z_t "introduces" an additional jump. By Markov property, after this jump, repeat "killed stable with extra jump instead of killing"

→ Suppose $Z_t^* = X_t^* \overset{\sim}{\sum} \varphi(x \leftarrow t)$

$$\overset{\sim}{\sum}_t = \overset{\sim}{\sum}_t^L + \overset{\sim}{\sum}_t^c$$

where $\overset{\sim}{\sum}_t^L$ is the same Lévy process as $\overset{\sim}{\sum}_t^*$ but without killing i.e. $\mathbb{E}[e^{-\vartheta \overset{\sim}{\sum}_t^L}] = e^{-(\Phi^*(\vartheta) - \vartheta^*)t}$

$\overset{\sim}{\sum}_t^c$ is a process which jumps at the times of a Poisson process with rate q^*
 (hopefully $\overset{\sim}{\sum}_t^c$ is a CPP process!)

has to be a CPP otherwise $\overset{\sim}{\sum}_t$ is not a L.p.
 and $\overset{\sim}{\sum}_t$ has to be a L.p. because the Lamperti transform says so!
 Intuitively speaking $\overset{\sim}{\sum}_t^c \perp \overset{\sim}{\sum}_t^L$

From previous lecture:

$$\Psi^*(\theta) = \frac{\Gamma(\alpha - i\theta)}{\Gamma(\alpha \hat{r} - i\theta)} \times \frac{\Gamma(1 + i\theta)}{\Gamma(1 - \alpha \hat{r} + i\theta)}$$

$$q^* = \Psi^*(0) = \frac{\Gamma(\alpha)}{\Gamma(\alpha \hat{r}) \Gamma(1 - \alpha \hat{r})}$$

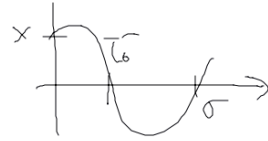
$$\rightarrow \underline{LS}(\beta, \gamma, \hat{r}) = (1 - \alpha \hat{r}, \alpha \hat{r}, \alpha \hat{r})$$

For full description of $\Psi(\theta) := -\log \mathbb{E}[e^{i\theta \sum_{t=1}^{\infty} X_t}]$
we need to work out jump distⁿ of $\sum_{t=1}^{\infty} X_t$.

Lemma: For each $x > 0$,

the joint law of $(X_{\tau_0^-}, X_{\tau_0^-}, X_\sigma)$ under \mathbb{P}_x is equal to $(x X_{\tau_0^-}, x X_{\tau_0^-}, x X_\sigma)$ under \mathbb{P}_1

where $\sigma = \inf\{t > \tau_0^-, X_t > 0\}$.



Pf $\forall c > 0, X_t^c = c X_{c^{-\alpha}t}, t \geq 0.$

$$(X^c, \mathbb{P}_x) \stackrel{d}{=} (X, \mathbb{P}_{cx})$$

$$\tau_0^{c^-} = \inf\{t > 0 : X_t^c < 0\}$$

$$\sigma^c = \inf\{t > \tau_0^{c^-} : X_t^c > 0\}$$

$$c^\alpha \tau_0^- = c^\alpha \inf\{t > 0 : X_{tc^{-\alpha}} < 0\}$$

$$= \inf\{s > 0 : c X_{sc^{-\alpha}} < 0\}$$

$$= \tau_0^{c^-}$$

Similarly $c^\alpha \sigma = \sigma^c$

$$(c X_{\tau_0^-}, c X_{\tau_0^-}, c X_\sigma)$$

$$= (c X_{c^{-\alpha} \tau_0^{c^-}}, c X_{c^{-\alpha} \tau_0^{c^-}}, c X_{c^{-\alpha} \sigma^c})$$

$$\stackrel{d}{=} (X_{\tau_0^-}, X_{\tau_0^-}, X_\sigma)$$

Let $x = \frac{1}{c}$

What are the jumps of \sum^c ?

$$\sum_t^c = \sum_t^L \oplus \sum_t^c$$

$$\Delta Z_{T_0^-} = Z_{T_0^-} - Z_{T_0^-}$$

$$T_0^- = \inf\{t > 0 : X_t < 0\}$$

$$= T_1 = \inf\{t > 0 : \Delta \sum_t^c \neq 0\}$$

$$\begin{aligned} \Delta Z_{T_0^-} &= e^{\sum_{T_1}^c} - e^{\sum_{T_1^-}^c} && \text{under } \mathbb{P}_1 \\ &= e^{\sum_{T_1^-}^c} \left(e^{\Delta \sum_{T_1}^c} - 1 \right) \\ &= X_{T_0^-} \left(e^{\Delta \sum_{T_1}^c} - 1 \right) \end{aligned}$$

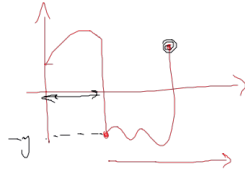
$$\begin{aligned} e^{\Delta \sum_{T_1}^c} &= 1 + \frac{\Delta Z_{T_0^-}}{X_{T_0^-}} = 1 + \frac{(X_{T_0} - X_{T_0^-})}{X_{T_0^-}} \\ &= \frac{X_{T_0}}{X_{T_0^-}} \end{aligned}$$

Note: take $0 < t_1 < t_2 < \dots < t_n \leq t$

$$\begin{aligned} & \mathbb{E}_1 \left[f(X_{t_1}, \dots, X_{t_n}) \mathbb{1}(t < T_0^-) g\left(\frac{X_t}{X_{T_0^-}}\right) \right] \\ & \stackrel{\text{under } \mathbb{P}_1}{=} \mathbb{E}_1 \left[f(X_{t_1}, \dots, X_{t_n}) \mathbb{1}(t < T_0^-) \mathbb{E}_{X_t} \left[g\left(\frac{X_t}{X_{T_0^-}}\right) \right] \right] \\ & \stackrel{\text{leave}}{=} \mathbb{E}_1 \left[f(X_{t_1}, \dots, X_{t_n}) \mathbb{1}(t < T_0^-) \right] \mathbb{E}_1 \left[g\left(\frac{X_t}{X_{T_0^-}}\right) \right] \end{aligned}$$

This shows independence of jumps of \sum^c with

Lemma The r.v. $e^{\Delta \tilde{S}_T^c}$
 is equal in distⁿ to $\left(-\frac{X_{\tau_0^-}}{X_{\tau_0^-}}\right) \times \underline{\hat{X}_{\hat{\tau}_0^-}}$
 where $\hat{X} \perp\!\!\!\perp X$, $\hat{X} \stackrel{d}{=} -X$ and
 $\hat{\tau}_0^- = \inf\{t > 0 : \hat{X}_t < 0\}$.



From previously $e^{\Delta \tilde{S}_T^c} = \frac{X_\sigma}{X_{\tau_0^-}}$

$$\mathbb{E}_1\left[f\left(\frac{X_\sigma}{X_{\tau_0^-}}\right) \mid \mathcal{F}_{\tau_0^-}\right]$$

$$= \mathbb{E}_{-y}\left[f\left(-\hat{X}_{\hat{\tau}_0^-}\right)\right] \Big|_{y=X_{\tau_0^-}}$$

Lemma

$$= \mathbb{E}_1\left[f\left(y \hat{X}_{\hat{\tau}_0^-}\right)\right] \Big|_{y=X_{\tau_0^-}}$$

$$\Rightarrow \mathbb{P}_1\left(X_\sigma \in \cdot \mid \mathcal{F}_{\tau_0^-}\right) = \mathbb{P}_1\left(y \hat{X}_{\hat{\tau}_0^-} \in \cdot \mid y=X_{\tau_0^-}\right)$$

Now compute

$$\mathbb{E}_1\left[f\left(\frac{X_\sigma}{X_{\tau_0^-}}\right)\right] = \mathbb{E}_1\left[\int_{(0,\infty)} f\left(\frac{x}{X_{\tau_0^-}}\right) \mathbb{P}_1(X_\sigma \in dx \mid \mathcal{F}_{\tau_0^-})\right]$$

$$= \mathbb{E}_1\left[\int_{(0,\infty)} f\left(\frac{x}{X_{\tau_0^-}}\right) \mathbb{P}_1(y \hat{X}_{\hat{\tau}_0^-} \in dx) \Big|_{y=X_{\tau_0^-}}\right]$$

$$= \mathbb{E}_1\left[\mathbb{E}_1\left[f\left(\frac{y \hat{X}_{\hat{\tau}_0^-}}{z}\right) \Big|_{y=X_{\tau_0^-}, z=X_{\tau_0^-}}\right]\right]$$

$$\Rightarrow \frac{X_\sigma}{X_{\tau_0^-}} \stackrel{d}{=} \left(\frac{X_{\tau_0^-}}{X_{\tau_0^-}}\right) \otimes \hat{X}_{\hat{\tau}_0^-}$$

Note $\mathbb{E}_c(\theta) = \int_0^x (1-e^{-\theta y}) F_c(dx)$
 $= \int_0^x (1 - \mathbb{E}[e^{i\theta \Delta \tilde{S}_1^c}])$

$$\mathbb{E}[e^{i\theta \Delta \tilde{S}_1^c}] = \mathbb{E}\left[\left(\frac{X_{T_0^-}}{X_{T_0^+}}\right)^{i\theta}\right] \times \mathbb{E}\left[(-\hat{X}_{T_0^+})^{i\theta}\right]$$

Recall:

$$\mathbb{P}(X_{T_0^+} - a \in du, a - X_{T_0^-} \in dv, a - \bar{X}_{T_0^+} \in dy)$$

$$= \frac{\sin \pi \alpha}{\pi} \frac{\Gamma(\alpha+1)}{\Gamma(\alpha)\Gamma(\alpha\beta)} \frac{(a-y)^{\alpha-1} (v-y)^{\alpha\beta-1}}{(v+u)^{1+\alpha}}$$

$$\mathbb{E}_1\left[(-\hat{X}_{T_0^+})^{i\theta}\right] = \mathbb{E}\left[(X_{T_0^+} - 1)^{i\theta}\right]$$

$$= \frac{\sin \pi \alpha}{\pi} \int_0^\infty t^{i\theta - \alpha} (1+t)^{-1} dt$$

$$= \frac{\sin \pi \alpha}{\pi} \frac{\Gamma(1-\alpha+i\theta)\Gamma(\alpha-i\theta)}{\Gamma(\alpha)}$$

$$\mathbb{E}\left[\left(\frac{X_{T_0^-}}{X_{T_0^+}}\right)^{i\theta}\right] = \mathbb{E}\left[\frac{(\hat{X}_{T_0^+} - 1)^{i\theta}}{(1 - \hat{X}_{T_0^+})^{i\theta}}\right]$$

$$= \frac{\sin \pi \alpha}{\pi} \frac{\Gamma(\alpha+1)}{\Gamma(\alpha)\Gamma(\alpha\beta)}$$

$$\times \int_0^1 \int_0^\infty \int_0^\infty \frac{u^{i\theta} (1-y)^{\alpha\beta-1} (v-y)^{\alpha\beta-1}}{v^{i\theta} (v+u)^{1+\alpha}} du dv dy$$

$$\stackrel{\text{lemma}}{=} \frac{\Gamma(i\theta+1)\Gamma(\alpha-i\theta)}{\Gamma(\alpha)}$$

In conclusion: lemma, $\theta \in \mathbb{R}$

$$\mathbb{E}[e^{i\theta \Delta \tilde{S}_1^c}] = \frac{\sin \pi \alpha}{\pi} \frac{\Gamma(1-\alpha+i\theta)\Gamma(\alpha-i\theta)}{\Gamma(\alpha)}$$

$$\times \Gamma(1+i\theta)\Gamma(\alpha-i\theta)$$

with lemma!

$$\mathbb{P}(\Delta \tilde{S}_1^c \in dx) = \frac{\alpha \Gamma(\alpha+1)\Gamma(\alpha\beta+1)}{\Gamma(\alpha)\Gamma(1-\alpha)\Gamma(\alpha+2)}$$

$$\times e^{-\alpha x} {}_2F_1(1, \alpha+1; \alpha+2; 1-e^{-x})$$

(need to understand ${}_2F_1(a, b; c; x)$ in the sense of analytic continuation)

Putting the pieces together

$$\begin{aligned} \tilde{\Psi}(\theta) &= \Psi_L(\theta) + \Psi_C(\theta) \\ &= (\Psi^*(\theta) - q^*) + q^* (1 - \mathbb{E}[e^{i\theta \Delta \sum_{t=1}^n \tilde{S}_t}]) \\ &= \Psi^*(\theta) - q^* \mathbb{E}[e^{i\theta \Delta \sum_{t=1}^n \tilde{S}_t}] \end{aligned}$$

$$= \frac{\Gamma(\alpha - i\theta) \Gamma(1 + i\theta)}{\Gamma(\alpha \hat{\rho} - i\theta) \Gamma(1 - \alpha \hat{\rho} + i\theta)} \left[\frac{1}{\Gamma(\alpha \rho) \Gamma(1 - \alpha \rho)} \frac{\Gamma(1 - \alpha \rho + i\theta) \Gamma(\alpha \rho - i\theta) \Gamma(\alpha - i\theta)}{\Gamma(\alpha)} \right]$$

$$= \Gamma(\alpha - i\theta) \Gamma(1 + i\theta) \frac{\Gamma(1 - \alpha \rho + i\theta) \Gamma(\alpha \rho - i\theta)}{\Gamma(\alpha \hat{\rho} - i\theta) \Gamma(1 - \alpha \hat{\rho} + i\theta) \Gamma(\alpha \rho) \Gamma(1 - \alpha \rho)}$$

$$= \Gamma(\alpha - i\theta) \Gamma(1 + i\theta) \Gamma(1 - \alpha \rho + i\theta) \Gamma(\alpha \rho - i\theta) \times \left[\frac{\Im(\alpha \hat{\rho} - i\theta) \Im(\alpha \rho - i\theta)}{\pi^2} - \frac{\Im(\alpha \hat{\rho} \pi) \Im(\alpha \rho \pi)}{\pi^2} \right]$$

$$= \frac{\Im(\pi(\alpha \hat{\rho} - i\theta)) \Im(\pi(\alpha \rho - i\theta)) + \Im(\pi i \theta) \Im(\pi(\alpha - i\theta))}{\pi^2}$$

$$= \Im(\pi \alpha \hat{\rho}) \Im(\pi \alpha \rho)$$

$$= \frac{\Gamma(\alpha - i\theta) \Gamma(1 + i\theta) \Gamma(1 - \alpha \rho + i\theta) \Gamma(\alpha \rho - i\theta)}{\Gamma(\alpha \hat{\rho} - i\theta) \Gamma(1 - \alpha \hat{\rho} + i\theta) \Gamma(\alpha \rho) \Gamma(1 - \alpha \rho)} \Gamma(1 + z) = z \Gamma(z)$$

$$= \frac{\Gamma(\alpha \rho - i\theta) \Gamma(1 - \alpha \rho + i\theta)}{\Gamma(-i\theta) \Gamma(1 - \alpha + i\theta)}$$

$$\in \mathcal{H}_G(\beta, \delta, \hat{\beta}, \hat{\delta})$$

$$= (1, \alpha \rho, 1 - \alpha, \alpha \hat{\rho}) \in \mathcal{A}_3$$

$$\text{if } \alpha \in (1, 2)$$

$$\in \mathcal{A}_1 \text{ if } \alpha \in (0, 1]$$